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Nonlinear Prediction, Filtering and Smoothing

10.1 Introduction
The Kalman filter is widely used for linear estimation problems where its behaviour is well-understood. Under prescribed conditions, the estimated states are unbiased and stability is guaranteed. Many real-world problems are nonlinear which requires amendments to linear solutions. If the nonlinear models can be expressed in a state-space setting then the Kalman filter may find utility by applying linearisations at each time step. Linearising means finding tangents to the curves of interest about the current estimates, so that the standard filter recursions can be employed in tandem to produce predictions for the next step. This approach is known as extended Kalman filtering – see [1] – [5].

Extended Kalman filters (EKFs) revert to optimal Kalman filters when the problems become linear. Thus, EKFs can yield approximate minimum-variance estimates. However, there are no accompanying performance guarantees and they fall into the try-at-your-own-risk category. Indeed, Anderson and Moore [3] caution that the EKF “can be satisfactory on occasions”. A number of compounding factors can cause performance degradation. The approximate linearisations may be crude and are carried out about estimated states (as opposed to true states). Observability problems occur when the variables do not map onto each other, giving rise to discontinuities within estimated state trajectories. Singularities within functions can result in non-positive solutions to the design Riccati equations and lead to instabilities.

The discussion includes suggestions for performance improvement and is organised as follows. The next section begins with Taylor series expansions, which are prerequisites for linearisation. First, second and third-order EKFs are then derived. EKFs tend be prone to instability and a way of enforcing stability is to masquerade the design Riccati equation by a faux version. This faux algebraic Riccati equation technique [6] – [10] is presented in Section 10.3. In Section 10.4, the higher order terms discarded by an EKF are treated as uncertainties. It is shown that a robust EKF arises by solving a scaled $H_{\infty}$ problem in lieu of one possessing uncertainties. Nonlinear smoother procedures can be designed similarly. The use of fixed-lag and Rauch-Tung-Striebel smoothers may be preferable from a complexity perspective. However, the approximate minimum-variance and robust smoothers, which are presented in Section 10.5, revert to optimal solutions when the nonlinearities and uncertainties diminish. Another way of guaranteeing stability is to by imposing constraints and one such approach is discussed in Section 10.6.

“It is the mark of an instructed mind to rest satisfied with the degree of precision to which the nature of the subject admits and not to seek exactness when only an approximation of the truth is possible.”

Aristotle

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10.2 Extended Kalman Filtering

10.2.1 Taylor Series Expansion

A nonlinear function \( a_i(x) : \mathbb{R}^n \to \mathbb{R} \) having \( n \) continuous derivatives may be expanded as a Taylor series about a point \( x_0 \)

\[
a_i(x) = a_i(x_0) + \frac{1}{1!}(x - x_0)^T \nabla a_i(x_0) + \frac{1}{2!}(x - x_0)^T \nabla^2 a_i(x_0)(x - x_0) + \cdots
\]

where \( \nabla a_i = \left[ \frac{\partial a_i}{\partial x_1}, \ldots, \frac{\partial a_i}{\partial x_n} \right] \) is known as the gradient of \( a_i() \) and

\[
\nabla^2 a_i = \left[ \begin{array}{cccc}
\frac{\partial^2 a_i}{\partial x_1^2} & \frac{\partial^2 a_i}{\partial x_1 \partial x_2} & \cdots & \frac{\partial^2 a_i}{\partial x_1 \partial x_n} \\
\frac{\partial^2 a_i}{\partial x_2 \partial x_1} & \frac{\partial^2 a_i}{\partial x_2^2} & \cdots & \frac{\partial^2 a_i}{\partial x_2 \partial x_n} \\
\vdots & \vdots & \ddots & \vdots \\
\frac{\partial^2 a_i}{\partial x_n \partial x_1} & \frac{\partial^2 a_i}{\partial x_n \partial x_2} & \cdots & \frac{\partial^2 a_i}{\partial x_n^2}
\end{array} \right]
\]

is called a Hessian matrix.

10.2.2 Nonlinear Signal Models

Consider nonlinear systems having state-space representations of the form

\[
x_{i+1} = a_i(x_i) + b_i(x_i)w_i, \tag{2}
\]

\[
y_i = c_i(x_i), \tag{3}
\]

where \( a_i(), b_i() \) and \( c_i() \) are continuous differentiable functions. For a scalar function, \( a_i(x) : \mathbb{R} \to \mathbb{R} \), its Taylor series about \( x = x_0 \) may be written as

“In the real world, nothing happens at the right place at the right time. It is the job of journalists and historians to correct that.” *Samuel Langhorne Clemens aka. Mark Twain*
Taylor series about a point $x_k$:

$$
a_k(x) = a_k(x_0) + (x - x_0) \frac{\partial a_k}{\partial x} \bigg|_{x=x_0} + \frac{1}{2}(x - x_0)^2 \frac{\partial^2 a_k}{\partial x^2} \bigg|_{x=x_0} + \frac{1}{6}(x - x_0)^3 \frac{\partial^3 a_k}{\partial x^3} \bigg|_{x=x_0} + \cdots + \frac{1}{n!}(x - x_0)^n \frac{\partial^n a_k}{\partial x^n} \bigg|_{x=x_0}.
$$

Similarly, Taylor series for $b_k(x) : \mathbb{R} \to \mathbb{R}$ and $c_k(x) : \mathbb{R} \to \mathbb{R}$ about $x = x_0$ are

$$
b_k(x) = b_k(x_0) + (x - x_0) \frac{\partial b_k}{\partial x} \bigg|_{x=x_0} + \frac{1}{2}(x - x_0)^2 \frac{\partial^2 b_k}{\partial x^2} \bigg|_{x=x_0} + \frac{1}{6}(x - x_0)^3 \frac{\partial^3 b_k}{\partial x^3} \bigg|_{x=x_0} + \cdots + \frac{1}{n!}(x - x_0)^n \frac{\partial^n b_k}{\partial x^n} \bigg|_{x=x_0},
$$

and

$$
c_k(x) = c_k(x_0) + (x - x_0) \frac{\partial c_k}{\partial x} \bigg|_{x=x_0} + \frac{1}{2}(x - x_0)^2 \frac{\partial^2 c_k}{\partial x^2} \bigg|_{x=x_0} + \frac{1}{6}(x - x_0)^3 \frac{\partial^3 c_k}{\partial x^3} \bigg|_{x=x_0} + \cdots + \frac{1}{n!}(x - x_0)^n \frac{\partial^n c_k}{\partial x^n} \bigg|_{x=x_0},
$$

respectively.

### 10.2.3 First-Order Extended Kalman Filter

Suppose that filtered estimates $\hat{x}_{k|k}$ of $x_k$ are desired given observations

$$
z_k = c_k(x_k) + v_k,
$$

where $v_k$ is a measurement noise sequence. A first-order EKF for the above problem is developed below. Following the approach within [3], the nonlinear system (2) – (3) is approximated by

$$
x_{k+1} = A_k x_k + B_k w_k + \mu_k,
$$

$$
y_k = C_k x_k + \pi_k,
$$

where $A_k$, $B_k$, $C_k$, $\mu_k$ and $\pi_k$ are found from suitable truncations of the Taylor series for each nonlinearity. From Chapter 4, a filter for the above model is given by

$$
\hat{x}_{k|k} = \hat{x}_{k|k-1} + L_k (z_k - C_k \hat{x}_{k|k-1} - \pi_k),
$$

$$
\hat{x}_{k+1|k} = A_k \hat{x}_{k|k} + \mu_k,
$$

"You will always define events in a manner which will validate your agreement with reality." *Steve Anthony Maraboli*
where $L_k = P_{k-1/k}C^T_k \Omega_k^{-1}$ is the filter gain, in which $\Omega_k = C_k P_{k-1/k}C_k^T + R_k$, $P_{k-1/k} = P_{k-1/k-1} - P_{k-1/k-1}C_k^T(C_k P_{k-1/k-1}C_k^T + R_k)^{-1}C_k P_{k-1/k-1}$ and $P_{k-1/k} = A_k P_{k-1/k}A_k^T + B_k Q_k B_k^T$. It is common practice (see [1] - [5]) to linearise about the current conditional mean estimate, retain up to first order terms within the corresponding Taylor series and assume $B_k = b_k(\hat{x}_{k|k})$. This leads to

$$a_k(x) = a_k(\hat{x}_{k|k}) + (x - \hat{x}_{k|k})^T \nabla a_k \bigg|_{x = \hat{x}_{k|k}} \mu_k$$

(12)

and

$$c_k(x) = c_k(\hat{x}_{k|k-1}) + (x - \hat{x}_{k|k-1})^T \nabla c_k \bigg|_{x = \hat{x}_{k|k-1}}$$

(13)

where $A_k = \nabla a_k \bigg|_{x = \hat{x}_{k|k}}$, $C_k = \nabla c_k \bigg|_{x = \hat{x}_{k|k}}$, $\mu_k = a_k(\hat{x}_{k|k}) - A_k \hat{x}_{k|k}$ and $\pi_k = c_k(\hat{x}_{k|k-1}) - C_k \hat{x}_{k|k-1}$. Substituting for $\mu_k$ and $\pi_k$ into (10) - (11) gives

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + L_k (z_k - c_k(\hat{x}_{k|k-1}))$$

(14)

$$\hat{x}_{k+1|k} = a_k(\hat{x}_{k|k})$$

(15)

Note that nonlinearities enter into the state correction (14) and prediction (15), whereas linearised matrices $A_k$, $B_k$ and $C_k$ are employed in the Riccati equation and gain calculations.

In the case of scalar states, the linearisations are $A_k = \frac{\partial a_k}{\partial x} \bigg|_{x = \hat{x}_{k|k}}$ and $C_k = \frac{\partial c_k}{\partial x} \bigg|_{x = \hat{x}_{k|k}}$. In texts on optimal filtering, the recursions (14) - (15) are either called a first-order EKF or simply an EKF, see [1] - [5]. Two higher order versions are developed below.

### 10.2.4 Second-Order Extended Kalman Filter

Truncating the series (1) after the second-order term and observing that $(x - \hat{x}_{k|k})^T \nabla \hat{x}_{k|k}$ is a scalar yields

$$a_k(x) = a_k(\hat{x}_{k|k}) + (x - \hat{x}_{k|k})^T \nabla a_k \bigg|_{x = \hat{x}_{k|k}} + \frac{1}{2} (x - \hat{x}_{k|k})^T \nabla^2 \nabla a_k \bigg|_{x = \hat{x}_{k|k}} (x - \hat{x}_{k|k})$$

$$= a_k(\hat{x}_{k|k}) + (x - \hat{x}_{k|k})^T \nabla a_k \bigg|_{x = \hat{x}_{k|k}} + \frac{1}{2} \nabla P_{k-1/k} \nabla^T a_k \bigg|_{x = \hat{x}_{k|k}}$$

(16)

"People take the longest possible paths, digress to numerous dead ends, and make all kinds of mistakes. Then historians come along and write summaries of this messy, nonlinear process and make it appear like a simple straight line." Dean L. Kamen

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where \( A_k = \nabla a_k(x) \big|_{x=x_{k|k}} \) and \( \mu_k = a_k(\hat{x}_{k|k}) - A_k\hat{x}_{k|k} + \frac{1}{2} \nabla P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}} \). Similarly for the system output,

\[
c_k(x) = c_k(\hat{x}_{k|k}) + (x_k - \hat{x}_{k|k})^T \nabla c_k \big|_{x=x_{k|k}} + \frac{1}{2} (x_k - \hat{x}_{k|k})^T \nabla^T \nabla c_k \big|_{x=x_{k|k}} (x_k - \hat{x}_{k|k})
\]

\[
= c_k(\hat{x}_{k|k}) + (x_k - \hat{x}_{k|k})^T \nabla c_k \big|_{x=x_{k|k}} + \frac{1}{2} \nabla P_{k|k} \nabla^T c_k \big|_{x=x_{k|k}}
\]

\[
= C_k x_k + n_k,
\]

where \( C_k = \nabla c_k(x) \big|_{x=x_{k|k}} \) and \( n_k = c_k(\hat{x}_{k|k}) - C_k\hat{x}_{k|k} + \frac{1}{2} \nabla P_{k|k} \nabla^T c_k \big|_{x=x_{k|k}} \). Substituting for \( \mu_k \) and \( n_k \) into the filtering and prediction recursions (10) – (11) yields the second-order EKF

\[
\hat{x}_{k|k} = \hat{x}_{k|k-1} + L_k \left( z_k - C_k(\hat{x}_{k|k-1}) - \frac{1}{2} \nabla P_{k|k-1} \nabla^T c_k \big|_{x=x_{k|k-1}} \right),
\]

\[
\hat{x}_{k|k-1} = a_k(\hat{x}_{k|k}) + \frac{1}{2} \nabla P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}}.
\]

The above form is described in [2]. The further simplifications \( \nabla P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}} \approx \text{tr} \left( \nabla P_{k|k} \nabla^T c_k \big|_{x=x_{k|k}} \right) \) and \( \nabla P_{k|k-1} \nabla^T c_k \big|_{x=x_{k|k-1}} \approx \text{tr} \left( \nabla P_{k|k-1} \nabla^T c_k \big|_{x=x_{k|k-1}} \right) \) are assumed in [4], [5].

### 10.2.5 Third-Order Extended Kalman Filter

Higher order EKFs can be realised just as elegantly as its predecessors. Retaining up to third-order terms within (1) results in

\[
a_k(x) = a_k(\hat{x}_{k|k}) + (x_k - \hat{x}_{k|k}) \nabla a_k \big|_{x=x_{k|k}} + \frac{1}{2} \nabla P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}}
\]

\[
+ \frac{1}{6} \nabla^2 P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}} (x_k - \hat{x}_{k|k}) \nabla a_k \big|_{x=x_{k|k}}
\]

\[
= A_k x_k + \mu_k,
\]

where

\[
A_k = \nabla a_k(x) \big|_{x=x_{k|k}} + \frac{1}{6} \nabla^2 P_{k|k} \nabla^T a_k \big|_{x=x_{k|k}}
\]

"It might be a good idea if the various countries of the world would occasionally swap history books, just to see what other people are doing with the same set of facts." William E. Vaughan
and \( \mu_k = a_k(\hat{x}_{k|k-1}) - A_k\hat{x}_{k|k-1} + \frac{1}{2} V P_{k|k-1} V^T a_k |_{x=x_{k|k-1}} \). Similarly, for the output nonlinearity it is assumed that

\[
c_k(x_k) \approx c_k(\hat{x}_{k|k-1}) + (x - \hat{x}_{k|k-1}) V c_k |_{x=x_{k|k-1}} + \frac{1}{6} P_{k|k-1} V^T c_k |_{x=x_{k|k-1}} + \frac{1}{6} V P_{k|k-1} V^T c_k |_{x=x_{k|k-1}} \]

\[= C_k x_k + \pi_k, \]

where

\[
C_k = V c_k |_{x=x_{k|k-1}} + \frac{1}{6} V P_{k|k-1} V^T c_k |_{x=x_{k|k-1}} \]

and \( \pi_k = c_k(\hat{x}_{k|k-1}) - C_k \hat{x}_{k|k-1} + \frac{1}{6} V P_{k|k-1} V^T c_k |_{x=x_{k|k-1}} \). The resulting third-order EKF is defined by (18) – (19) in which the gain is now calculated using (21) and (23).

**Example 1.** Consider a linear state evolution \( x_{k+1} = A x_k + w_k \) with \( A = 0.5, w_k \in \mathbb{R}, Q = 0.05, \) a nonlinear output mapping \( y_k = \sin(x_k) \) and noisy observations \( z_k = y_k + v_k, v_k \in \mathbb{R} \). The first-order EKF for this problem is given by

\[
\hat{x}_{k|k} = \hat{x}_{k|k-1} + L_k (z_k - \sin(\hat{x}_{k|k-1})) ,
\]

\[
\hat{x}_{k+1|k} = A \hat{x}_{k|k} ,
\]

where \( L_k = P_{k|k-1} C_k^T \Omega_k^{-1} \), \( \Omega_k = C_k P_{k|k-1} C_k^T + R_k \), \( C_k = \cos(\hat{x}_{k|k-1}) \), \( P_{k|k} = P_{k|k-1} - P_{k|k-1} C_k^T (C_k P_{k|k-1} C_k^T + R_k)^{-1} C_k P_{k|k-1} \) and \( R_{k+1} = A P_{k|k-1} A^T + Q_k \). The filtering step within the second-order EKF is amended to

\[
\hat{x}_{k+1|k} = \hat{x}_{k+1|k-1} + L_k (z_k - \sin(\hat{x}_{k|k-1}) + \sin(\hat{x}_{k|k-1}) P_{k|k-1} / 2) .
\]

The modified output linearisation for the third-order EKF is

\[
C_k = \cos(\hat{x}_{k|k-1}) + \sin(\hat{x}_{k|k-1}) P_{k|k-1} / 6 .
\]

Simulations were conducted in which the signal-to-noise-ratio was varied from 20 dB to 40 dB for \( N = 200,000 \) realisations of Gaussian noise sequences. The mean-square-errors exhibited by the first, second and third-order EKFs are plotted in Fig. 1. The figure demonstrates that including higher-order Taylor series terms within the filter can provide small performance improvements but the benefit diminishes with increasing measurement noise.
Improvements but the benefit diminishes with increasing measurement noise.

The first, second and third-order EKFs are plotted in Fig. 1. The figure demonstrates that

by (18) – (19) in which the gain is now calculated using (21) and (23).

The modified output linearisation for the third-order EKF is nonlinear output mapping

for second-order EKF is amended to

\[
\begin{align*}
P_{C|C} & = \frac{1}{1} \\text{for} \quad N = 200,000 \text{ realisations of Gaussian noise sequences.}
\end{align*}
\]

The mean-square-errors exhibited by \( \pi \)

\[
\begin{align*}
\hat{\mu}_k & = \frac{1}{1} \\text{for} \quad N = 200,000 \text{ realisations of Gaussian noise sequences.}
\end{align*}
\]

Consider a linear state evolution

A nonlinear state evolution together with a nonlinear output mapping

where \( c_0(.) \) are assumed to be continuous differentiable functions.

Suppose that it is desired to calculate estimates of the states from the measurements. A nonlinear observer may be constructed having the form

\[
\hat{x}_{k+1|k} = A\hat{x}_k + g_0(z_k - c(\hat{x}_{k|k-1}))
\]

where \( g_0(.) \) is a gain function to be designed. From (24) – (26), the state prediction error is given by

"The observer, when he seems to himself to be observing a stone, is really, if physics is to be believed, observing the effects of the stone upon himself." Bertrand Arthur William Russell

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\[
\tilde{x}_{t+1|k} = A \tilde{x}_{t|k-1} - g_t(c_t) + w_t,
\]

(27)

where \( \tilde{x}_t = x_k - \hat{x}_{t+1|k-1} \) and \( \tilde{e}_t = z_k - c(\hat{x}_{t+1|k-1}) \). The Taylor series expansion of \( c(.) \) to first order terms leads to \( \tilde{e}_t \approx C_t \tilde{x}_{t|k-1} + v_t \), where \( C_t = \nabla c_t(x)|_{x=\hat{x}_{t+1|k-1}} \). The objective here is to design \( g_t(\tilde{e}_t) \) to be a linear function of \( \tilde{x}_{t|k-1} \) to first order terms. It will be shown that for certain classes of problems, this objective can be achieved by a suitable choice of a linear bounded matrix function of the states \( D_k \) resulting in the time-varying gain function \( g_t(\tilde{e}_t) = K_t D_k \tilde{e}_k \) where \( K_t \) is a gain matrix of appropriate dimension. For example, consider \( n_t \in \mathbb{R}^n \) and \( z_k \in \mathbb{R}^m \), which yield \( \tilde{e}_t \in \mathbb{R}^n \) and \( C_t \in \mathbb{R}^{mn} \). Suppose that a linearisation \( D_t \in \mathbb{R}^{mn} \) can be found so that \( C_t = D_t C_k \in \mathbb{R}^{mn} \) possesses approximately constant terms. Then the locally linearised error (27) may be written as

\[
\tilde{x}_{t+1|k} = (A - K_t \bar{C}_t) \tilde{x}_{t|k-1} - K_t D_t v_t + w_t.
\]

(28)

If \( |\lambda_i(A)| < 1, i = 1 \ldots n_t \) and if the pair \( (A, \bar{C}_t) \) is completely observable, then the asymptotic stability of (28) can be guaranteed by selecting the gain such that \( |\lambda_i(A - K_t \bar{C}_t)| < 1 \). A method for selecting the gain is described below.

### 10.3.2 Gain Selection

From (28), an approximate equation for the error covariance \( P_{t+1|k} = E(\tilde{x}_{t+1|k-1}) \) is

\[
P_{t+1|k} = (A - K_t \bar{C}_t) P_{t|k-1} (A - K_t \bar{C}_t)^T + K_t D_t R_t K_t^T + Q,
\]

(29)

which can be written as

\[
P_{t+1|k} = P_{t|k-1} - P_{t|k-1} \bar{C}_t^T (\bar{C}_t P_{t|k-1} \bar{C}_t^T + D_t R_t D_t^T)^{-1} \bar{C}_t P_{t|k-1} + Q,
\]

(30)

\[
P_{t+1|k} = AP_{t|k-1} A^T + Q.
\]

(31)

In an EKF for the above problem, the gain is obtained by solving the above Riccati difference equation and calculating

\[
K_t = P_{t|k-1} \bar{C}_t^T (\bar{C}_t P_{t|k-1} \bar{C}_t^T + D_t R_t D_t^T)^{-1}.
\]

(32)

The faux algebraic Riccati equation approach [6] - [10] is motivated by connections between Riccati difference equation and algebraic Riccati equation solutions. Indeed, it is noted for some nonlinear problems that the gains can converge to a steady-state matrix [3]. This technique is also known as ‘covariance setting’. Following the approach of [10], the Riccati difference equation (30) may be masqueraded by the faux algebraic Riccati equation

\[
\Sigma_t = \Sigma_t - \Sigma_t \bar{C}_t^T (\bar{C}_t \Sigma_t \bar{C}_t^T + D_t R_t D_t^T)^{-1} \bar{C}_t \Sigma_t.
\]

(33)

"The universe as we know it is a joint product of the observer and the observed." Pierre Teilhard De Chardin

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That is, rather than solve (30), an arbitrary positive definite solution $\Sigma_k$ is assumed instead and then the gain at each time $k$ is calculated from (31) – (32) using $\Sigma_k$ in place of $P_{k/k-1}$.

10.3.3 Tracking Multiple Signals

Consider the problem of tracking two frequency or phase modulated signals which may be modelled by equation (34), where $a_k^{(1)}, a_k^{(2)}, \alpha_k^{(1)}, \alpha_k^{(2)}, \phi_k^{(1)}, \phi_k^{(2)}, \mu_k^{(1)}, \mu_k^{(2)}, \mu_\nu^{(1)} \mu_\nu^{(2)} \in \mathbb{R}$ and $w_k^{(1)}, \ldots, w_k^{(6)} \in \mathbb{R}$ are zero-mean, uncorrelated, white processes with covariance $Q = \text{diag}(\sigma_{\omega_1}^2, \ldots, \sigma_{\omega_2}^2)$. The states $a_k^{(1)}, a_k^{(2)}$ and $\phi_k^{(1)}, \phi_k^{(2)}, i = 1, 2,$ represent the signals’ instantaneous amplitude, frequency and phase components, respectively.

$$
\begin{align*}
\begin{bmatrix}
a_{k+1}^{(1)} \\
\alpha_{k+1}^{(1)} \\
\phi_{k+1}^{(1)} \\
a_{k+1}^{(2)} \\
\alpha_{k+1}^{(2)} \\
\phi_{k+1}^{(2)}
\end{bmatrix} &=
\begin{bmatrix}
\mu_k^{(1)} \\
0 \\
0 \\
\mu_k^{(2)} \\
0 \\
0
\end{bmatrix}
\begin{bmatrix}
a_k^{(1)} \\
0 \\
0 \\
a_k^{(2)} \\
0 \\
0
\end{bmatrix}
+ 
\begin{bmatrix}
w_k^{(1)} \\
w_k^{(2)} \\
w_k^{(3)} \\
w_k^{(4)} \\
w_k^{(5)} \\
w_k^{(6)}
\end{bmatrix}.
\end{align*}
$$

(34)

Let

$$
\begin{align*}
\begin{bmatrix}
z_k^{(1)} \\
z_k^{(2)} \\
z_k^{(3)} \\
z_k^{(4)}
\end{bmatrix} &=
\begin{bmatrix}
ad_k^{(1)} \cos \phi_k^{(1)} \\
ad_k^{(1)} \sin \phi_k^{(1)} \\
ad_k^{(2)} \cos \phi_k^{(2)} \\
ad_k^{(2)} \sin \phi_k^{(2)}
\end{bmatrix}
+ 
\begin{bmatrix}
v_k^{(1)} \\
v_k^{(2)} \\
v_k^{(3)} \\
v_k^{(4)}
\end{bmatrix}.
\end{align*}
$$

(35)

denote the complex baseband observations, where $v_k^{(1)}, \ldots, v_k^{(4)} \in \mathbb{R}$ are zero-mean, uncorrelated, white processes with covariance $R = \text{diag}(\sigma_{\nu_1}^2, \ldots, \sigma_{\nu_2}^2)$. Expanding the prediction error to linear terms yields $C_k = [C_k^{(1)} C_k^{(2)}]$, where

$$
C_k^{(1)} =
\begin{bmatrix}
\cos \phi_k^{(1)} & 0 & -\bar{a}_k^{(1)} \cos \phi_k^{(1)} \\
\sin \phi_k^{(1)} & 0 & \bar{a}_k^{(1)} \sin \phi_k^{(1)}
\end{bmatrix}.
$$

This form suggests the choice $D_k = \begin{bmatrix} D_k^{(1)} & D_k^{(2)} \end{bmatrix}$, where

$$
D_k^{(1)} =
\begin{bmatrix}
\cos \phi_k^{(1)} & \sin \phi_k^{(1)} \\
-\sin \phi_k^{(1)} & \cos \phi_k^{(1)}
\end{bmatrix}.
$$

"If you haven’t found something strange during the day, it hasn’t been much of a day.” John Archibald Wheeler
In the multiple signal case, the linearization $\sum_i = D_i C_i$ does not result in perfect decoupling. While the diagonal blocks reduce to 
\[
\sum_{i=1}^{n} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0 & 1 \end{bmatrix},
\]
the off-diagonal blocks are
\[
\sum_{i=1}^{n} = \begin{bmatrix} \cos(\phi_i - \phi_i) & 0 \\ -\frac{1}{\sin(\phi_i - \phi_i)} & \cos(\phi_i - \phi_i) \end{bmatrix}.
\]
Assuming a symmetric positive definite solution to (33) of the form $\Sigma = \begin{bmatrix} \Sigma & 0 & 0 \\ 0 & \Sigma & \Sigma \end{bmatrix}$, where $\Sigma$, $\Sigma$, $\Sigma$, $\Sigma \in \mathbb{R}$ and choosing the gains according to (32) yields $K = \begin{bmatrix} K' \\ 0 \\ 0 \end{bmatrix}$, where $K' = \Sigma(\Sigma + \sigma)^{-1}$, $K' = \Sigma(\Sigma + \sigma)^{-1}$ and $K' = \Sigma(\Sigma + \sigma)^{-1}$. The nonlinear observer then becomes
\[
\begin{align*}
\dot{\tilde{u}}_i &= \tilde{u}_i(1) + \Sigma_i (z_i(1) \cos(z_i(2) + z_i(1) \sin(z_i(2))))(\Sigma + \sigma)^{-1}, \\
\dot{\tilde{a}}_i &= \tilde{a}_i(1) + \Sigma_i (z_i(1) \cos(z_i(2) + z_i(1) \sin(z_i(2))))(\Sigma + \sigma)^{-1}, \\
\dot{\tilde{\phi}}_i &= \tilde{\phi}_i(1) + \Sigma_i (z_i(1) \cos(z_i(2) + z_i(1) \sin(z_i(2))))(\Sigma + \sigma)^{-1}.
\end{align*}
\]

### 10.3.4 Stability Conditions
In order to establish conditions for the error system (28) to be asymptotically stable, the problem is recast in a passivity framework as follows. Let $w = \begin{bmatrix} u_1 \\ u_2 \\ \vdots \\ u_n \end{bmatrix}$, $e = \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix} \in \mathbb{R}^n$.

Consider the configuration of Fig. 2, in which there is a cascade of a stable linear system $G$ and a nonlinear function matrix $\gamma$ acting on $e$. It follows from the figure that
\[
e = w - G \gamma(e).
\]

Let $\Delta$ denote a forward difference operator with $\Delta e_i = e_i - e_i[-1]$. It is assumed that $\gamma$ satisfies some sector conditions which may be interpreted as bounds existing on the slope of the components of $\gamma$; see Theorem 14, p. 7 of [11].

"Discovery consists of seeing what everyone has seen and thinking what nobody has thought." Albert Szent-Györgyi

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Let us consider the configuration of Fig. 2, in which there is a cascade of a stable linear system with the components of a non-linear observer then becomes 

\[ \Delta_k \] denotes a forward difference operator with 

\[ f_k \] acting on \( \gamma \) consisting of non-interacting nonlinearities, with \( \gamma(e^{(i)}) \) monotonically increasing in the sector \([0, \beta], \beta \geq 0, \beta \in \mathbb{R}, \) \( \) is a causal, stable, finite-gain, time-invariant map \( \mathbb{R}^n \rightarrow \mathbb{R}^n, \) having a z-transform \( G(z) \), which is bounded on the unit circle. Let \( I \) denote an \( m \times m \) identity matrix. Suppose that for some \( q > 0, q \in \mathbb{R}, \) there exists a \( \delta \in \mathbb{R}, \) such that 

\[ \langle (G(z) + q\Delta G(z) + 1\beta^{-1})e, e \rangle \geq \delta \langle e, e \rangle \]  

(38) 

for all \( e^{(i)} \in \mathbb{R}, \) \( e^{(i)} \neq 0. \) Assume that \( \gamma(e) \in \ell_2. \)

**Proof:** From (36), \( \Delta_e w = \Delta_e + \Delta G(z) \gamma(e) \) and \( w + q\Delta_e w = (G(z) + q\Delta G(z) + 1\beta^{-1}) \gamma(e) + e - 1\beta^{-1} \gamma(e) + e - 1\beta^{-1} \gamma(e) + q\Delta_e e. \) Then 

\[ \langle w + q\Delta_e w, \gamma(e) \rangle \geq \langle e - 1\beta^{-1} \gamma(e), \gamma(e) \rangle + \langle q\Delta_e e, \gamma(e) \rangle \]

+ \( \langle (G(z) + q\Delta G(z) + 1\beta^{-1}) \gamma(e), \gamma(e) \rangle. \)  

(39)

Consider the first term on the right hand side of (39). Since the \( \gamma(e) \) consists of non-interacting nonlinearities, \( \gamma(e), \gamma(e) = \sum_{i=1}^{m} \gamma(e^{(i)}), e^{(i)} \) and \( \langle e - 1\beta^{-1} \gamma(e), \gamma(e) \rangle = \sum_{i=1}^{m} \langle e^{(i)} - \beta^{-1} \gamma(e^{(i)}), e^{(i)} \rangle \geq 0. \) Using the approach of [11] together with the sector conditions on the identical non-interacting nonlinearities (37), it can be shown that expanding out the second term of (39) yields 

"The intelligent man finds almost everything ridiculous, the sensible man hardly anything." Johann Wolfgang von Goethe
Using \( \|\lambda v\| \leq 2 \|\nu\| \) (from p. 192 of [11]), the Schwartz inequality and the triangle inequality, it can be shown that
\[
\langle w + qv, \gamma(\nu) \rangle \leq (1 + 2q) \|w\|.
\] (40)

It follows from (38) – (40) that \( \|\gamma(\nu)\| \leq (1 + 2q) \|\delta\| \); hence \( \gamma(\nu) \in \ell_2 \). Since the gain of \( G(z) \) is finite, it also follows that \( G(z) \gamma(\nu) \in \ell_2 \).

If \( G(z) \) is stable and bounded on the unit circle, then the test condition (38) becomes
\[
\lambda_{\text{min}} [I + q(I - z^{-1})G(z) + G^*(z) + \beta^2] \geq \delta,
\] (41)
see pp. 175 and 194 of [11].

### 10.3.5 Applications

**Example 2** [10]. Consider a unity-amplitude frequency modulated (FM) signal modelled as
\[
\dot{\phi}_{k+1} = \mu_\nu \phi_k + w_\nu, \quad \phi_{k+1} = \phi_k + \omega_\nu, \quad z^{(1)}_k = \cos(\phi_k) + v^{(1)}_k \quad \text{and} \quad z^{(2)}_k = \sin(\phi_k) + v^{(2)}_k.
\]
The error system for an FM demodulator may be written as
\[
\begin{bmatrix}
\dot{\phi}_{k+1} \\
\dot{\phi}_{k+1}
\end{bmatrix} =
\begin{bmatrix}
\mu_\nu & 0 \\
1 & 1
\end{bmatrix}
\begin{bmatrix}
\phi_k \\
\phi_k
\end{bmatrix}
+ \begin{bmatrix} K_1 \\ K_2 \end{bmatrix} \sin(\phi_k) + w_k
\] (42)
for gains \( K_1, K_2 \in \mathbb{R} \) to be designed. In view of the form (36), the above error system is reformatted as
\[
\begin{bmatrix}
\dot{\phi}_{k+1} \\
\dot{\phi}_{k+1}
\end{bmatrix} =
\begin{bmatrix}
\mu_\nu & -K_1 \\
1 & 1-K_2
\end{bmatrix}
\begin{bmatrix}
\phi_k \\
\phi_k
\end{bmatrix}
+ \begin{bmatrix} K_1 \\ K_2 \end{bmatrix} \sin(\phi_k) + w_k
\] (43)
where \( \gamma(x) = x - \sin(x) \). The \( z \)-transform of the linear part of (43) is
\[
G(z) = (K_2z + K_2 + \mu_\nu K_1) \quad (z^2 + (K_2 - 1 - \mu_\nu)z + K_1 + 1 - \mu_\nu K_2)^{-1}.
\]
The nonlinearity satisfies the sector condition (37) for \( \beta = 1.22 \). Candidate gains may be assessed by checking that \( G(z) \) is stable and the test condition (41). The stable gain space calculated for the case of \( \mu_\nu = 0.9 \) is plotted in Fig. 3. The gains are required to lie within the shaded region of the plot for the error system (42) to be asymptotically stable.
A speech utterance, namely, the phrase “Matlab is number one”, was sampled at 8 kHz and used to synthesize a unity-amplitude FM signal. An EKF demodulator was constructed for the above model with \( \sigma_w^2 = 0.02 \). In a nonlinear observer design it was found that suitable parameter choices were \( \Sigma_k = \begin{bmatrix} 0.001 & 0.08 \\ 0.08 & 0.7 \end{bmatrix} \). The nonlinear observer gains were censored at each time \( k \) according to the stable gain space of Fig. 3. The results of a simulation study using 100 realisations of Gaussian measurement noise sequences are shown in Fig. 4. The figure demonstrates that enforcing stability can be beneficial at low SNR, at the cost of degraded high-SNR performance.

**Example 3** [10]. Suppose that there are two superimposed FM signals present in the same frequency channel. Neglecting observation noise, a suitable approximation of the demodulator error system in the form (36) is given by

\[
\begin{bmatrix}
\hat{\phi}_1^{(1)} \\
\hat{\phi}_1^{(2)} \\
\hat{\phi}_2^{(1)} \\
\hat{\phi}_2^{(2)}
\end{bmatrix} = (A - K_2 \bar{C})
\begin{bmatrix}
\hat{\phi}_1^{(0)} \\
\hat{\phi}_1^{(0)} \\
\hat{\phi}_2^{(0)} \\
\hat{\phi}_2^{(0)}
\end{bmatrix} + K_1
\begin{bmatrix}
\sin(\hat{\phi}_1^{(0)}) - \hat{\phi}_1^{(0)} \\
\sin(\hat{\phi}_2^{(0)}) - \hat{\phi}_2^{(0)}
\end{bmatrix},
\]

where \( A = \text{diag}(A^{(1)}, A^{(2)}) \), \( A^{(1)} = \begin{bmatrix} \mu_\omega & 0 \\ 1 & 1 \end{bmatrix} \), \( \bar{C} = \begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \). The linear part of (44) may be written as \( G(z) = \bar{C}(zI - (A - K_2 \bar{C}))^{-1}K_1 \). Two 8-kHz speech utterances, “Matlab is number one” and “Number one is Matlab”, centred at ±0.25 rad/s, were used to synthesize two superimposed unity-amplitude FM signals. Simulations were conducted using 100 realisations of Gaussian measurement noise sequences. The test condition (41) was

“To avoid criticism, do nothing, say nothing, be nothing.” Elbert Hubbard
evaluated at each time $k$ for the above parameter values with $\beta = 1.2$, $q = 0.001$, $\delta = 0.82$ and used to censor the gains. The resulting co-channel demodulation performance is shown in Fig. 5. It can be seen that the nonlinear observer significantly outperforms the EKF at high SNR.

Two mechanisms have been observed for occurrence of outliers or faults within the co-channel demodulators. Firstly errors can occur in the state attribution, that is, there is correct tracking of some component speech message segments but the tracks are inconsistently associated with the individual signals. This is illustrated by the example frequency estimate tracks shown in Figs. 6 and 7. The solid and dashed lines in the figures indicate two sample co-channel frequency tracks. Secondly, the phase unwrapping can be erroneous so that the frequency tracks bear no resemblance to the underlying messages. These faults can occur without any significant deterioration in the error residual.

![Figure 6](image1.png)  ![Figure 7](image2.png)

Figure 6. Sample EKF frequency tracks for Example 3.  Figure 7. Sample Nonlinear observer frequency tracks for Example 3.

The EKF demodulator is observed to be increasingly fault prone at higher SNR. This arises because lower SNR designs possess narrower bandwidths and so are less sensitive to nearby frequency components. The figures also illustrate the trade-off between stability and optimality. In particular, it can be seen from Fig. 6, that the sample EKF speech estimates exhibit faults in the state attribution. This contrasts with Fig. 7, where the nonlinear observer’s estimates exhibit stable state attribution at the cost of degraded speech fidelity.

10.4 Robust Extended Kalman Filtering

10.4.1 Nonlinear Problem Statement

Consider again the nonlinear, discrete-time signal model (2), (7). It is shown below that the $H_k$ techniques of Chapter 9 can be used to recast nonlinear filtering problems into a model uncertainty setting. The following discussion attends to state estimation, that is, $C_{1,k} = I$ is assumed within the problem and solution presented in Section 9.3.2.

“You have enemies? Good. That means you’ve stood up for something, sometime in your life.” *Winston Churchill*
The Taylor series expansions of the nonlinear functions \( a_i(\cdot) \), \( b_i(\cdot) \) and \( c_i(\cdot) \) about filtered and predicted estimates \( \hat{x}_{i,k} \) and \( \hat{x}_{i,k-1} \) may be written as

\[
a_i(x_k) = a_i(\hat{x}_{i,k}) + \nabla a_i(\hat{x}_{i,k}) (x_k - \hat{x}_{i,k}) + \Delta_i(\hat{x}_{i,k}) ,
\]

\[
b_i(x_k) = b_i(\hat{x}_{i,k}) + \Delta_i(\hat{x}_{i,k}) ,
\]

\[
c_i(x_k) = c_i(\hat{x}_{i,k-1}) + \nabla c_i(\hat{x}_{i,k-1}) (x_k - \hat{x}_{i,k-1}) + \Delta_i(\hat{x}_{i,k-1}) ,
\]

where \( \Delta_i(\cdot) , \Delta_2(\cdot) , \Delta_3(\cdot) \) are uncertainties that account for the higher order terms, \( \hat{x}_{i,k} = x_k - \hat{x}_{i,k} \) and \( \hat{x}_{i,k-1} = x_k - \hat{x}_{i,k-1} \). It is assumed that \( \Delta_i(\cdot) , \Delta_2(\cdot) \) and \( \Delta_3(\cdot) \) are continuous operators mapping \( \ell_2 \to \ell_2 \), with \( H_\infty \) norms bounded by \( \delta_1 , \delta_2 \) and \( \delta_3 \) respectively.

Substituting (45) – (47) into the nonlinear system (2), (7) gives the linearised system

\[
x_{i,k} = A_i x_k + B_i w_k + \mu_k + \Delta_i(\hat{x}_{i,k}) + \Delta_2(\hat{x}_{i,k}) w_k ,
\]

\[
z_k = C_i x_k + \pi_k + \Delta_i(\hat{x}_{i,k-1}) + \nu_k ,
\]

where \( A_i = \nabla a_i(x) \big|_{x=\hat{x}_{i,k-1}}, C_i = \nabla c_i(x) \big|_{x=\hat{x}_{i,k-1}}, \mu_k = a_i(\hat{x}_{i,k}) - A_i \hat{x}_{i,k} \) and \( \pi_k = c_i(\hat{x}_{i,k-1}) - C_i \hat{x}_{i,k-1} \).

Note that the first-order EKF for the above system arises by setting the uncertainties \( \Delta_i(\cdot) , \Delta_2(\cdot) \) and \( \Delta_3(\cdot) \) to zero as

\[
\hat{x}_{i,k} = \hat{x}_{i,k-1} + L_k (z_k - c_i(\hat{x}_{i,k-1})) ,
\]

\[
\hat{x}_{i,k-1} = a_i(\hat{x}_{i,k}) ,
\]

\[
L_k = P_{i,k-1} C_i^T (C_i P_{i,k-1} C_i^T + R_k)^{-1} ,
\]

\[
P_{i,k} = P_{i,k-1} - P_{i,k-1} C_i^T (C_i P_{i,k-1} C_i^T + R_k)^{-1} P_{i,k-1} C_i ,
\]

\[
P_{i,k-1} = A_i P_{i,k-1} A_i^T + B_i Q_i B_i^T .
\]

### 10.4.2 Robust Solution

Following the approach in Chapter 9, instead of addressing the problem (48) – (49) which possesses uncertainties, an auxiliary \( H_\infty \) problem is defined as

\[
x_{i,k+1} = A_i x_k + B_i w_k + \mu_k + s_k ,
\]

\[
z_k = C_i x_k + \pi_k + \nu_k + f_k ,
\]

“Fight the good fight.” Timothy 4:7
\[ \hat{x}_{k+1} = x_k - \hat{x}_{k+1}, \]  

where \( s_k = \Delta_t(\hat{x}_{k+1}) + \Delta_t(\hat{x}_{k+1})w_k \) and \( t_k = \Delta_t(\hat{x}_{k+1}) \approx \Delta_t(\hat{x}_{k+1}) \) are additional exogenous inputs satisfying

\[
\|s_k\| \leq \delta^2 \|\hat{x}_{k+1}\| + \delta^2 \|w_k\|, \tag{58} 
\]

\[
\|t_k\| \leq \delta^2 \|\hat{x}_{k+1}\| + \delta^2 \|\hat{x}_{k+1}\|. \tag{59} 
\]

A sufficient solution to the auxiliary \( H_e \) problem (55) – (57) can be obtained by solving another problem in which \( w_k \) and \( t_k \) are scaled in lieu of the additional inputs \( s_k \) and \( t_k \). The scaled \( H_e \) problem is defined by

\[
x_{r+1} = A_rx_k + B_riw_k + \mu_k, \tag{60} 
\]

\[ z_k = C_0x_k + C_wv_k + \pi_k, \tag{61} \]

\[ \hat{x}_{r+1} = x_k - \hat{x}_{r+1}, \tag{62} \]

where \( c_w, c_r \in \mathbb{R} \) are to be found.

**Lemma 2 [12]:** The solution of the \( H_e \) problem (60) – (62), where \( v_k \) is scaled by

\[ c_v^2 = 1 - \gamma^2 \delta_1^2 - \gamma^2 \delta_2^2, \tag{63} \]

and \( w_k \) is scaled by

\[ c_w^2 = c_v^2 (1 + \delta_1^2)^{-1}, \tag{64} \]

is sufficient for the solution of the auxiliary \( H_e \) problem (55) – (57).

**Proof:** If the \( H_e \) problem (50) – (52) has been solved then there exists a \( \gamma \neq 0 \) such that

\[
\|s_k\| \leq \gamma^2 (\|w_k\| + \|t_k\| + \|v_k\|),
\]

\[
\leq \gamma^2 (\|w_k\| + \delta^2 \|\hat{x}_{k+1}\| + \delta^2 \|\hat{x}_{k+1}\| + \delta^2 \|w_k\| + \delta^2 \|\hat{x}_{k+1}\| + \|v_k\|),
\]

which implies

\[
(1 - \gamma^2 \delta_1^2 - \gamma^2 \delta_2^2) \|s_k\| \leq \gamma^2 ((1 + \delta_2^2) \|w_k\| + \|v_k\|)
\]

“You can’t wait for inspiration. You have to go after it with a club.” *Jack London*

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Lemma 2 \[12\]: The solution of the H system is sufficient for the solution of another problem in which the auxiliary H is scaled by \(w_k\) where \(s_t = \gamma^2(c^2 \| w_k \|^2 + c^2 \| e_k \|^2)\).

The robust first-order extended Kalman filter for state estimation is given by (50) – (52),

\[
P_{k|k} = P_{k|k-1} - \left[ I - C_k P_{k|k-1} C_k^T \right] \left[ P_{k|k-1} - \gamma^2 I - P_{k|k-1} C_k^T \right]^{-1} \left[ P_{k|k-1} C_k C_k^T \right] P_{k|k-1}
\]

and (54). As discussed in Chapter 9, a search is required for a minimum \(\gamma\) such that

\[
\left[ P_{k|k-1} - \gamma^2 I - P_{k|k-1} C_k^T \right] > 0 \quad \text{and} \quad P_{k|k} > 0 \quad \text{over} \quad k \in [1, N].
\]

An illustration is provided below.

Figure 8. Histogram of demodulator mean-square-error for Example 4: (i) first-order EKF (solid line) and first-order robust EKF (dotted line).

Example 4 [12]. Suppose that an FM signal is generated by

\[
\begin{align*}
\phi_{k+1} &= \phi_k + \omega_k, \\
\omega_{k+1} &= \mu_\omega \omega_k + w_\omega, \\
\phi_{k+1}^{(1)} &= \arctan(\mu_\phi \phi_k + \phi_k), \\
z_k^{(1)} &= \cos(\phi_k^{(1)}) + v_k^{(1)}, \\
z_k^{(2)} &= \sin(\phi_k^{(1)}) + v_k^{(2)}.
\end{align*}
\]

The objective is to construct an FM demodulator that produces estimates of the frequency message \(\phi_k\) from the noisy in-phase and quadrature measurements \(z_k^{(1)}\) and \(z_k^{(2)}\), respectively. Simulations were conducted with \(\mu_\omega = 0.9, \mu_\phi = 0.99\) and \(\sigma^2_{z^{(1)}} = \sigma^2_{z^{(2)}} = 0.001\). It was found for \(\sigma^2_w < 0.1\), where the state behaviour is almost linear, a robust EKF does not

“Happy is he who gets to know the reasons for things.” Virgil
improve on the EKF. However, when $\sigma^2_\omega = 1$, the problem is substantially nonlinear and a performance benefit can be observed. A robust EKF demodulator was designed with

$$x_{k+1} = \begin{bmatrix} \phi_k \\ \omega_k \end{bmatrix}, A_k = \begin{bmatrix} \mu_y & (\mu_y \phi_k + \omega_k)^2 + 1 \\ 0 & (\mu_y \phi_k + \omega_k)^2 + 1 \end{bmatrix}, C_k = \begin{bmatrix} -\sin(\phi_{k+1}) \\ \cos(\phi_{k+1}) \end{bmatrix},$$

$\delta_1 = 0.1, \delta_2 = 4.5$ and $\delta_3 = 0.001$. It was found that $\gamma = 1.38$ was sufficient for $P_{k/k-1}$ of the above Riccati difference equation to always be positive definite. A histogram of the observed frequency estimation error is shown in Fig. 8, which demonstrates that the robust demodulator provides improved mean-square-error performance. For sufficiently large $\sigma^2_\omega$, the output of the above model will resemble a digital signal, in which case a detector may outperform a demodulator.

### 10.5 Nonlinear Smoothing

#### 10.5.1 Approximate Minimum-Variance Smoother

Consider again a nonlinear estimation problem where $x_{k+1} = a_k(x_k) + B_k w_k, z_k = c_k(x_k) + v_k$, with $x_k \in \mathbb{R}$, in which the nonlinearities $a_k(\cdot), c_k(\cdot)$ are assumed to be smooth, differentiable functions of appropriate dimension. The linearizations akin to Extended Kalman filtering may be applied within the smoothers described in Chapter 7 in the pursuit of performance improvement. The fixed-lag, Fraser-Potter and Rauch-Tung-Striebel smoother recursions are easier to apply as they are less complex. The application of the minimum-variance smoother can yield approximately optimal estimates when the problem becomes linear, provided that the underlying assumptions are correct.

**Procedure 1.** An approximate minimum-variance smoother for output estimation can be implemented via the following three-step procedure.

**Step 1.** Operate

$$a_k = -\Omega_k^{1/2}(z_k - c_k(\hat{x}_{k|k-1})), \tag{69}$$

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + L_k(z_k - c_k(\hat{x}_{k|k-1})), \tag{70}$$

$$\tilde{x}_{k|k} = a_k(\hat{x}_{k|k}), \tag{71}$$

on the measurement $z_k$, where $L_k = P_{k|k-1}C_k^T \Omega_k^{-1}$,

$$\Omega_k = C_kP_{k|k-1}C_k^T + R_k,$$

$$P_{k|k} = P_{k|k-1} - P_{k|k-1}C_k^T \Omega_k^{-1}C_k P_{k|k-1},$$

$$P_{k+1|k} = A_kP_{k|k}A_k^T + B_kQ_k B_k^T,$$

$$A_k = \frac{\partial a_k}{\partial x} \bigg|_{x = \tilde{x}_{k|k}} \quad \text{and} \quad C_k = \frac{\partial c_k}{\partial x} \bigg|_{x = \hat{x}_{k|k-1}}.$$

“You can recognize a pioneer by the arrows in his back” Beverly Rubik

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Step 2. Operate (69) – (71) on the time-reversed transpose of \( a_k \). Then take the time-reversed transpose of the result to obtain \( \beta_k \).

Step 3. Calculate the smoothed output estimate from

\[ \hat{y}_{k,N} = z_k - R_k \beta_k. \]

### 10.5.2 Robust Smoother

From the arguments within Chapter 9, a smoother that is robust to uncertain \( w_k \) and \( \nu_k \) can be realised by replacing the error covariance correction (72) by

\[ P_{k+1} = P_{k+1} - P_{k+1} \left[ C_k \right]^T C_k \left[ C_k P_{k+1} C_k^T - \gamma^2 I - C_k P_{k+1} C_k^T \gamma^2 \right]^{-1} C_k \left[ C_k P_{k+1} C_k^T - \gamma^2 I + C_k P_{k+1} C_k^T \gamma^2 \right]^{-1} P_{k+1} \]

within Procedure 1. As discussed in Chapter 9, a search for a minimum \( \gamma \) such that

\[ C_k P_{k+1} C_k^T - \gamma^2 I - C_k P_{k+1} C_k^T \gamma^2 > 0 \] and \( P_{k+1} > 0 \) over \( k \in [1, N] \) is desired.

### 10.5.3 Application

Returning to the problem of demodulating a unity-amplitude FM signal, let

\[ A = \begin{bmatrix} \mu_x & 0 \\ 1 & \mu_x \end{bmatrix}, \quad B = \begin{bmatrix} 1 \\ 0 \end{bmatrix}, \quad z_k^{(0)} = \cos(\phi_k) + \nu_k^{(0)}, \quad z_k^{(2)} = \sin(\phi_k) + \nu_k^{(2)}, \]

where \( \omega_k, \phi_k, z_k \) and \( \nu_k \) denote the instantaneous frequency message, instantaneous phase, complex observations and measurement noise respectively. A zero-mean voiced speech utterance "a e i o u" was sampled at 8 kHz, for which estimates \( \hat{\mu}_x = 0.97 \) and \( \hat{\sigma}_w^2 = 0.053 \) were obtained using an expectation maximization algorithm. An FM discriminator output [13],

\[ z_k^{(3)} = \left( z_k^{(0)} \frac{dz_k^{(2)}}{dt} - z_k^{(2)} \frac{dz_k^{(0)}}{dt} \right) \left( \left( z_k^{(2)} \right)^2 + \left( z_k^{(0)} \right)^2 \right)^{-1}, \]

serves as a benchmark and as an auxiliary frequency measurement for the above smoother.

The innovations within Steps 1 and 2 are given by

\[
\begin{bmatrix}
\hat{z}_k^{(0)} \\
\hat{z}_k^{(2)} \\
\hat{z}_k^{(3)}
\end{bmatrix}
= \begin{bmatrix}
\cos(\hat{\xi}_k^{(2)}) \\
\sin(\hat{\xi}_k^{(2)}) \\
\hat{\xi}_k^{(2)}
\end{bmatrix}
and
\[
\begin{bmatrix}
\hat{a}_k^{(0)} \\
\hat{a}_k^{(2)} \\
\hat{a}_k^{(3)}
\end{bmatrix}
= \begin{bmatrix}
\cos(\hat{\xi}_k^{(3)}) \\
\sin(\hat{\xi}_k^{(3)}) \\
\hat{\xi}_k^{(3)}
\end{bmatrix},
\]

respectively. A unity-amplitude FM signal was synthesized using \( \mu_x = 0.99 \) and the SNR was varied in 1.5 dB steps from 3 dB to 15 dB. The mean-square errors were calculated over 200 realisations of Gaussian measurement noise and are shown in Fig. 9. It can be seen from the figure, that at 7.5 dB SNR, the first-order EKF improves on the FM discriminator MSE by about 12 dB. The improvement arises because the EKF

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"The farther the experiment is from theory, the closer it is to the Nobel Prize." Irène Joliot-Curie
demodulator exploits the signal model whereas the FM discriminator does not. The figure shows that the approximate minimum-variance smoother further reduces the MSE by about 2 dB, which illustrates the advantage of exploiting all the data in the time interval. In the robust designs, searches for minimum values of \( \gamma \) were conducted such that the corresponding Riccati difference equation solutions were positive definite over each noise realisation. It can be seen from the figure at 7.5 dB SNR that the robust EKF provides about a 1 dB performance improvement compared to the EKF, whereas the approximate minimum-variance smoother and the robust smoother performance are indistinguishable.

This nonlinear example illustrates once again that smoothers can outperform filters. Since a first-order speech model is used and the Taylor series are truncated after the first-order terms, some model uncertainty is present, and so the robust designs demonstrate a marginal improvement over the EKF.

![Graph showing FM demodulation performance comparison](image)

**Figure 9.** FM demodulation performance comparison: (i) FM discriminator (crosses), (ii) first-order EKF (dotted line), (iii) Robust EKF (dashed line), (iv) approximate minimum-variance smoother and robust smoother (solid line).

### 10.6 Constrained Filtering and Smoothing

#### 10.6.1 Background

Constraints often appear within navigation problems. For example, vehicle trajectories are typically constrained by road, tunnel and bridge boundaries. Similarly, indoor pedestrian trajectories are constrained by walls and doors. However, as constraints are not easily described within state-space frameworks, many techniques for constrained filtering and smoothing are reported in the literature. An early technique for constrained filtering involves augmenting the measurement vector with perfect observations [14]. The application of the perfect-measurement approach to filtering and fixed-interval smoothing is described in [15].

“They thought I was crazy, absolutely mad.” *Barbara McClintock*
Constraints can be applied to state estimates, see [16], where a positivity constraint is used within a Kalman filter and a fixed-lag smoother. Three different state equality constraint approaches, namely, maximum-probability, mean-square and projection methods are described in [17]. Under prescribed conditions, the perfect-measurement and projection approaches are equivalent [5], [18], which is identical to applying linear constraints within a form of recursive least squares.

In the state equality constrained methods [5], [16] – [18], a constrained estimate can be calculated from a Kalman filter’s unconstrained estimate at each time step. Constraint information could also be embedded within nonlinear models for use with EKFs. A simpler, low-computation-cost technique that avoids EKF stability problems and suits real-time implementation is described in [19]. In particular, an on-line procedure is proposed that involves using nonlinear functions to censor the measurements and subsequently applying the minimum-variance filter recursions. An off-line procedure for retrospective analyses is also described, where the minimum-variance fixed-interval smoother recursions are applied to the censored measurements. In contrast to the afore-mentioned techniques, which employ constraint matrices and vectors, here constraint information is represented by an exogenous input process. This approach uses the Bounded Real Lemma which enables the nonlinearities to be designed so that the filtered and smoothed estimates satisfy a performance criterion.

10.6.2 Problem Statement

The ensuing discussion concerns odd and even functions which are defined as follows. A function \( g_e \) of \( X \) is said to be odd if \( g_e(-X) = -g_e(X) \). A function \( f_e \) of \( X \) is said to be even if \( f_e(-X) = f_e(X) \). The product of \( g_e \) and \( f_e \) is an odd function since \( g_e(-X) f_e(-X) = -g_e(X) f_e(X) \).

Problems are considered where stochastic random variables are subjected to inequality constraints. Therefore, nonlinear censoring functions are introduced whose outputs are constrained to lie within prescribed bounds. Let \( \beta \in \mathbb{R}^r \) and \( g_e: \mathbb{R}^r \rightarrow \mathbb{R}^r \) denote a constraint vector and an odd function of a random variable \( X \in \mathbb{R}^r \) about its expected value \( E[X] \), respectively. Define the censoring function

\[
g(X) = E[X] + g_e(X, \beta), \tag{75}
\]

where

\[
g_e(X, \beta) = \begin{cases} 
\beta & \text{if } \beta \leq X - E[X] \\
X - E[X] & \text{if } -\beta < X - E[X] < \beta \\
-\beta & \text{if } X - E[X] \leq -\beta 
\end{cases} \tag{76}
\]

“If at first, the idea is not absurd, then there is no hope for it.” Albert Einstein
By inspection of (75) – (76), \( g(X) \) is constrained within \( E[X] \pm \beta \). Suppose that the probability density function of \( X \) about \( E[X] \) is even, that is, is symmetric about \( E[X] \). Under these conditions, the expected value of \( g(X) \) is given by

\[
E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx = E(X)\int_{-\infty}^{\infty} f(x)dx + \int_{-\infty}^{\infty} g_s(x, \beta)f_s(x)dx = E(X),
\]

(77)

since \( \int_{-\infty}^{\infty} f_s(x)dx = 1 \) and the product \( g_s(x, \beta)f_s(x) \) is odd.

Thus, a constraining process can be modelled by a nonlinear function. Equation (77) states that \( g(X) \) is unbiased, provided that \( g_s(X, \beta) \) and \( f_s(X) \) are odd and even functions about \( E[X] \), respectively. In the analysis and examples that follow, attention is confined to systems having zero-mean inputs, states and outputs, in which case the censoring functions are also centred on zero, that is, \( E[X] = 0 \).

Let \( w_k = [w_{1, k} \ldots w_{m, k}]^T \in \mathbb{R}^m \) represent a stochastic white input process having an even probability density function, with \( E[w_i] = 0 \), \( E[w_i w_i^T] = Q, \delta_{ij} \), in which \( \delta_{ij} \) denotes the Kronecker delta function. Suppose that the states of a system \( G: \mathbb{R}^m \rightarrow \mathbb{R}^r \) are realised by

\[
x_{k+1} = A_kx_k + B_kw_k, \quad (78)
\]

where \( A_k \in \mathbb{R}^{m \times m} \) and \( B_k \in \mathbb{R}^{m \times n} \). Since \( w_k \) is zero-mean, it follows that linear combinations of the states are also zero-mean. Suppose also that the system outputs, \( y_{ik} \), are generated by

\[
y_k = \begin{bmatrix} y_{1,k} \\ \vdots \\ y_{p,k} \end{bmatrix} = \begin{bmatrix} g_s(C_{i,1}x_1, \theta_{1,k}) \\ \vdots \\ g_s(C_{i,p}x_p, \theta_{p,k}) \end{bmatrix},
\]

(79)

where \( C_{i,j} \) is the \( j^{th} \) row of \( C \in \mathbb{R}^{n \times m} \), \( \theta_k = [\theta_{1,k} \ldots \theta_{p,k}]^T \in \mathbb{R}^p \) is an input constraint process and \( g_s(C_{i,j}x_1, \theta_{j,k}) \), \( j = 1, \ldots p \), is an odd censoring function centred on zero. The outputs \( y_{ijk} \) are constrained to lie within \( \pm \theta_{j,k} \), that is,

\[
-\theta_{j,k} \leq y_{ijk} \leq \theta_{j,k}. \quad (80)
\]

For example, if the system outputs represent the trajectories of pedestrians within a building then the constraint process could include knowledge about wall, floor and ceiling positions.

"It was not easy for a person brought up in the ways of classical thermodynamics to come around to the idea that gain of entropy eventually is nothing more nor less than loss of information." Gilbert Newton Lewis
Similarly, a vehicle trajectory constraint process could include information about building and road boundaries.

Assume that observations $z_k = y_k + v_k$ are available, where $v_k \in \mathbb{R}^r$ is a stochastic, white measurement noise process having an even probability density function, with $E[v_k] = 0$, $E[v_k v_k^T] = R, \delta_{j,k}$ and $E[w_k w_k^T] = 0$. It is convenient to define the stacked vectors $y = [y_1^T \ldots y_n^T]^T$ and $\theta = [\theta_1^T \ldots \theta_n^T]^T$. It follows that

$$\|y\|^2 \leq \|\theta\|^2.$$  

(81)

Thus, the energy of the system’s output is bounded from above by the energy of the constraint process.

The minimum-variance filter and smoother which produce estimates of a linear system’s output, minimise the mean square error. Here, it is desired to calculate estimates that trade off minimum mean-square-error performance and achieve

$$\|y\|^2 \leq \|\theta\|^2.$$  

(82)

Note that (80) implies (81) but the converse is not true. Although estimates $\hat{y}_{j,k}$ of $y_{j,k}$ satisfying $-\theta_{j,k} \leq \hat{y}_{j,k} \leq \theta_{j,k}$ are desirable, the procedures described below only ensure that (82) is satisfied.

10.6.3 Constrained Filtering

A procedure is proposed in which a linear filter $\mathcal{H}: \mathbb{R}^r \to \mathbb{R}^r$ is used to calculate estimates $\hat{y}$ from zero-mean measurements $z_k$ that are constrained using an odd censoring function to obtain

$$z_k = \begin{bmatrix} z_{1,k} \\ \vdots \\ z_{p,k} \end{bmatrix} = \begin{bmatrix} \mathcal{g}_e(z_{1,k}, y^\top \theta_{1,k}) \\ \vdots \\ \mathcal{g}_e(z_{p,k}, y^\top \theta_{p,k}) \end{bmatrix},$$  

(83)

which satisfy

$$\|z\|^2 \leq \gamma^{-2}\|\theta\|^2.$$  

(84)

where $z = [z_1^T \ldots z_p^T]^T$, for a positive $\gamma \in \mathbb{R}$ to be designed. This design problem is depicted in Fig. 10.

"Man’s greatest asset is the unsettled mind." Isaac Asimov
Censoring the measurements is suggested as a low-implementation-cost approach to constrained filtering. Design constraints are sought for the measurement censoring functions so that the outputs of a subsequent filter satisfy the performance objective (82). The recursions akin to the minimum-variance filter are applied to calculate predicted and filtered state estimates from the constrained measurements \( \tilde{z}_k \) at time \( k \). That is, the output mapping \( C_k \) is retained within the linear filter design even though nonlinearities are present with (83). The probability density functions associated with \( w_k \) and \( v_k \) are even; the nonlinear functions within (79) and (83) are odd; and the filter is initialized with \( \hat{x}_{0/0} = E\{x_0\} \).

**Lemma 3 [19]:** In respect of the filter (85) – (87) which operates on the constrained measurements (83), suppose the following:

(i) the probability density functions associated with \( w_k \) and \( v_k \) are even;  
(ii) the nonlinear functions within (79) and (83) are odd; and  
(iii) the filter is initialized with \( \hat{x}_{0/0} = E\{x_0\} \).

Then the following applies:

(i) the predicted state estimates, \( \hat{x}_{k+1/k} \), are unbiased;  
(ii) the corrected state estimates, \( \hat{x}_{k+1/k} \), are unbiased; and  
(iii) the output estimates, \( \hat{y}_{k+1/k} \), are unbiased.

“A mind that is stretched by a new idea can never go back to its original dimensions.” Oliver Wendell Holmes
Proof: (i) Condition (iii) implies $E[\hat{x}_{1:k}] = 0$, which is the initialization step of an induction argument. It follows from (87) that

$$\hat{x}_{k+1:k} = (A_k - K_k C_k)\hat{x}_{k:k-1} + K_k (z_k - C_k x_k - v_k).$$  \hspace{1cm} (88)

Subtracting (88) from (78) gives

$$E[\hat{x}_{k+1:k}] = (A_k - K_k C_k)E[\hat{x}_{k:k-1}] + B_k E[v_k] - K_k E[v_k] - K_k E[z_k - C_k x_k - v_k].$$

From above assumptions, the second and third terms on the right-hand-side of (89) are zero. The property (77) implies $E[z_k] = E[C_k x_k + v_k]$ and so $E[z_k - C_k x_k - v_k]$ is zero. The first term on the right-hand-side of (89) pertains to the unconstrained Kalman filter and is zero by induction. Thus, $E[\hat{x}_{k+1:k}] = 0$.

(ii) Condition (iii) again serves as an induction assumption. It follows from (86) that

$$\hat{x}_{k:k} = \hat{x}_{k:k-1} + L_k (C_k x_k + v_k - C_k \hat{x}_{k:k-1}) + L_k (z_k - C_k x_k - v_k).$$  \hspace{1cm} (90)

Substituting $x_k = A_k x_{k-1} + B_k w_{k-1}$ into (90) yields

$$\hat{x}_{k:k} = (I - L_k C_k)A_k \hat{x}_{k:k-1} + (I - L_k C_k)B_k w_{k-1} - L_k v_k - L_k (z_k - C_k x_k - v_k)$$

and $E[\hat{x}_{k:k}] = (I - L_k C_k)A_k E[\hat{x}_{k:k-1}] = (I - L_k C_k)A_k \ldots (I - L_k C_k)A_k E[\hat{x}_{0:k}]$. Hence, $E[\hat{x}_{k:k}] = 0$ by induction.

(iii) Defining $\tilde{y}_{k:k} = y_k - \hat{y}_{k:k} = C_k (x_k - \hat{x}_{k:k}) - C_k \hat{x}_{k:k} + y_k - C_k x_k$ and using (77) leads to $E[\tilde{y}_{k:k}] = C_k E[\hat{x}_{k:k}] + E[y_k - C_k x_k] = C_k E[\hat{x}_{k:k}] = 0$ under condition (iii).

Recall that the Bounded Real Lemma (see Lemma 7 of Chapter 9) specifies a bound for a system’s output and input energies. This lemma is used to find a design for $\gamma$ within (83) as described below.

Lemma 4 [19]: Consider the filter (85) – (87) which operates on the constrained measurements (83). Let $A_k = A_k - K_k C_k$, $B_k = K_k$, $C_k = C_k (I - L_k C_k)$ and $D_k = C_k L_k$ denote the state-space parameters of the filter. Suppose for a given $\gamma > 0$, that a solution $M_k = M_k^* > 0$ exists over $k \in [1, N]$ for the Riccati Difference equation resulting from the application of the Bounded Real Lemma to the system $\begin{bmatrix} A_k & B_k \\ C_k & D_k \end{bmatrix}$. Then the design $\gamma = \gamma_2$ within (83) results in the performance objective (82) being satisfied.

Proof: For the application of the Bounded Real Lemma to the filter (85) – (87), the existence of a solution $M_k = M_k^* > 0$ for the associated Riccati difference equation ensures that $\|\tilde{y}\|_2 \leq \gamma_2 \|\hat{x}\|_2$, $- x_k^T M_k x_k \leq \gamma_2^2 \|\hat{x}\|_2^2$, which together with (84) leads to (82).

It is argued below that the proposed filtering procedure is asymptotically stable.

“All truth passes through three stages: First, it is ridiculed; Second, it is violently opposed; and Third, it is accepted as self-evident.” Arthur Schopenhauer
Lemma 5 [19]: Define the filter output estimation error as \( \hat{y} = y - \hat{y} \). Under the conditions of Lemma 4, \( \hat{y} \in \ell_2 \).

Proof: It follows from \( \hat{y} = y - \hat{y} \) that \( \|\hat{y}\| \leq \|y\| + \|\hat{y}\| \), which together with (10) and the result of Lemma 4 yields \( \|\hat{y}\| \leq 2\|\hat{y}\| \), thus the claim follows.

10.6.4 Constrained Smoothing

In the sequel, it is proposed that the minimum-variance fixed-interval smoother recursions operate on the censored measurements \( z_k \) to produce output estimates \( \hat{y}_{k/N} \) of \( y_k \).

Lemma 6 [19]: In respect of the minimum-variance smoother recursions that operate on the censored measurements \( z_k \), under the conditions of Lemma 3, the smoothed estimates, \( \hat{y}_{k/N} \), are unbiased.

The proof follows mutatis mutandis from the approach within the proofs of Lemma 5 of Chapter 7 and Lemma 3. An analogous result to Lemma 5 is now stated.

Lemma 7 [19]: Define the smoother output estimation error as \( \hat{y} = y - \hat{y} \). Under the conditions of Lemma 3, \( \hat{y} \in \ell_2 \).

The proof follows mutatis mutandis from that of Lemma 5. Two illustrative examples are set out below. A GPS and inertial navigation system integration application is detailed in [19].

Example 5 [19]. Consider the saturating nonlinearity

\[
g_s(X, \beta) = 2\beta \pi^{-1} \arctan \left( \pi X (2\beta)^{-1} \right).
\]

which is a continuous approximation of (76) that satisfies \( |g_s(X, \beta)| \leq |\beta| \) and \( \frac{dg_s(X, \beta)}{dX} = \left( 1 + (\pi X)^2 (2\beta)^{-2} \right)^{-1} \approx 1 \) when \( (\pi X)^2 (2\beta)^{-2} \ll 1 \). Data was generated from (78), (79), (91), where

\[
A = \begin{bmatrix} 0.9 & 0 \\ 0 & 0.9 \end{bmatrix}, \quad B = C = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}, \quad \text{Gaussian, white, zero-mean processes with } Q = R = \begin{bmatrix} 0.01 & 0 \\ 0 & 0.01 \end{bmatrix}.
\]

The constraint vector within (80) was chosen to be fixed, namely, \( \beta_k = \begin{bmatrix} 0.5 \\ 0.5 \end{bmatrix} \),

\[k \in [1, 10^5].\]

The limits of the observed distribution of estimates, \( \hat{y}_{k/k} = \begin{bmatrix} \hat{y}_{1,k/k} \\ \hat{y}_{2,k/k} \end{bmatrix} \), arising by

“Everything we know is only some kind of approximation, because we know that we do not know all the laws yet. Therefore, things must be learned only to be unlearned again or, more likely, to be corrected.” Richard Phillips Feynman
operating the minimum-variance filter recursions on the raw data \( z_k = y_k + v_k \) are indicated by the outer black region of Fig. 11. It can be seen that the filter outputs do not satisfy the performance objective (82), which motivates the pursuit of constrained techniques. A minimum value of \( \gamma_2 = 1.24 \) was found for the solutions of the Riccati difference equation mentioned within Lemma 4 to be positive definite. The filter (85) – (87) was applied to the censored measurements \( \bar{z}_k = \begin{bmatrix} \bar{x}_k \\ \bar{x}_k \end{bmatrix} = \begin{bmatrix} g_z(z_{k+1}, \gamma \theta_{0,1}) \\ g_z(z_{k+1}, \gamma \theta_{0,1}) \end{bmatrix} \) using (91). The limits of the observed distribution of the constrained filter estimates are indicated by the inner white region of Fig. 11. The figure shows that the constrained filter estimates satisfy (82), which illustrates Lemma 5.

**Example 6 [19].** Measurements were similarly synthesized using the parameters of Example 5 to demonstrate constrained fixed-interval smoother performance. A minimum value of \( \gamma_2 = 5.6 \) was found for the solutions of the Riccati difference equation mentioned within Lemma 4 to be positive definite. The superimposed distributions of the unconstrained and constrained smoothers are respectively indicated by the inner and outer black regions of Fig. 12. It can be seen by inspection of the figure that the constrained smoother estimates meets (80), where as those produced by the standard smoother do not.

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Figure 11. Superimposed distributions of filtered estimates for Example 4: unconstrained filter (outer black); and constrained filter (middle white).

Figure 12. Superimposed distributions of smoothed estimates for Example 5: unconstrained smoother (outer black); and constrained smoother (middle white).

"An expert is a man who has made all the mistakes which can be made in a very narrow field." **Niels Henrik David Bohr**
The above examples involved searching for minimum value of $\gamma_2$ for the existence of positive definite solutions for the Riccati equation alluded to within Lemma 4. The need for a search may not be apparent as stability is guaranteed whenever a positive definite solution for the associated Riccati equation exists. Searching for a minimum $\gamma_2$ is advocated because the use of an excessively large value can lead to a nonlinearity design that is conservative and exhibits poor mean-square-error performance. If a design is still too conservative then an empirical value, namely, $\gamma_2 = \|s\|_2^{-1}$, may need to be considered instead.

**10.7 Conclusion**

In this chapter it is assumed that nonlinear systems are of the form $x_{k+1} = a_i(x_k) + b_i(u_k), y_k = c_i(x_k)$, where $a_i(\cdot), b_i(\cdot)$ and $c_i(\cdot)$ are continuous differentiable functions. The EKF arises by linearising the model about conditional mean estimates and applying the standard filter recursions. The first, second and third-order EKFs simplified for the case of $x_k \in \mathbb{R}$ are summarised in Table 1.

The EKF attempts to produce locally optimal estimates. However, it is not necessarily stable because the solutions of the underlying Riccati equations are not guaranteed to be positive definite. The faux algebraic Riccati technique trades off approximate optimality for stability. The familiar structure of the EKF is retained but stability is achieved by selecting a positive definite solution to a faux Riccati equation for the gain design.

$H_n$ techniques can be used to recast nonlinear filtering applications into a model uncertainty problem. It is demonstrated with the aid of an example that a robust EKF can reduce the mean square error when the problem is sufficiently nonlinear.

Linearised models may be applied within the previously-described smoothers in the pursuit of performance improvement. Nonlinear versions of the fixed-lag, Fraser-Potter and Rauch-Tung-Striebel smoothers are easier to implement as they are less complex. However, the application of the minimum-variance smoother can yield approximately optimal estimates when the problem becomes linear, provided that the underlying assumptions are correct. A smoother that is robust to input uncertainty is obtained by replacing the approximate error covariance correction with an $H_n$ version. The resulting robust nonlinear smoother can exhibit performance benefits when uncertainty is present.

In some applications, it may be possible to censor a system’s inputs, states or outputs, rather than proceed with an EKF design. It has been shown that the use of a nonlinear censoring function to constrain input measurements leads to bounded filter and smoother estimation errors.

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“Most of what I learned as an entrepreneur was by trial and error.” *Gordon Earl Moore*
The capacity to blunder slightly is the real marvel of DNA. Without this special attribute, we would still be anaerobic bacteria and there would be no music.” Lewis Thomas
Problem 3. Assuming the required derivatives exist.

(i) \( f(x) = a_0 + a_1(x-x_0) + a_2(x-x_0)^2 \).

(ii) \( f(x) = a_0 + a_1(x-x_0) + a_2(x-x_0)^2 + a_3(x-x_0)^3 \).

(iii) \( f(x,y) = a_0 + a_1(x-x_0) + a_2(x-x_0)^2 \) + \( a_3(y-y_0) + a_4(y-y_0)^2 + a_5(x-x_0)(y-y_0) \).

(iv) \( f(x,y) = a_0 + a_1(x-x_0) + a_2(x-x_0)^2 + a_3(x-x_0)^3 \) + \( a_4(y-y_0) + a_5(y-y_0)^2 + a_6(y-y_0)^3 \) + \( a_7(x-x_0)(y-y_0) + a_8(x-x_0)^2(y-y_0) \) + \( a_9(x-x_0)(y-y_0)^2 + a_{10}(x-x_0)^3(y-y_0) \) + \( a_{11}(x-x_0)(y-y_0)^3 + a_{12}(x-x_0)^2(y-y_0)^2 \).

(v) \( f(x,y) = a_0 + a_1(x-x_0) + a_2(x-x_0)^2 + a_3(x-x_0)^3 + a_4(x-x_0)^4 \) + \( a_5(y-y_0) + a_6(y-y_0)^2 + a_7(y-y_0)^3 + a_8(y-y_0)^4 \) + \( a_9(x-x_0)(y-y_0) + a_{10}(x-x_0)^2(y-y_0) \) + \( a_{11}(x-x_0)(y-y_0)^2 + a_{12}(x-x_0)^3(y-y_0) \) + \( a_{13}(x-x_0)(y-y_0)^3 + a_{14}(x-x_0)^2(y-y_0)^2 \).

Problem 2. Consider a state estimation problem, where \( x_{k+1} = a_0(x_k) + B_{w}w_{k}, y_{k} = c_{1}(x_{k}), z_{k} = y_{k} + v_{k} \), in which \( w_{k}, x_{0}, y_{0}, v_{0}, a_{1}, B_{w}, c_{1} \) \( \in \mathbb{R} \). Derive the

(i) first-order,

(ii) second-order,

(iii) third-order and

(iv) fourth-order EKFs,

assuming the required derivatives exist.

Problem 3. Suppose that an FM signal is generated by \( \phi_{k+1} = \phi_{k} + \omega_{k}, z_{k}^{(1)} = a_{0}\cos(\phi_{k}) + v_{k}^{(1)} \) and \( z_{k}^{(2)} = a_{0}\sin(\phi_{k}) + v_{k}^{(2)} \). Write down the recursions for

(i) first-order and

(ii) second-order

EKF demodulators.

"I am quite conscious that my speculations run quite beyond the bounds of true science." Charles Robert Darwin

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Problem 4. (Continuous-time EKF) Assume that continuous-time signals may be modelled as $\dot{x}(t) = a(x(t)) + w(t)$, $y(t) = c(x(t))$, $z(t) = y(t) + v(t)$, where $E[w(t)w^T(t)] = Q(t)$ and $E[v(t)v^T(t)] = R(t)$.

(i) Show that approximate state estimates can be obtained from $\dot{\hat{x}}(t) = a(\hat{x}(t)) + K(t)(z(t) - c(\hat{x}(t)))$, where $K(t) = P(t)C^T(t)R^{-1}(t)$, $\dot{P}(t) = A(t)P(t) + P(t)A^T(t) - K(t)C(t)P(t) + Q(t)$, and $\dot{C}(t) = \frac{\partial c(x)}{\partial x}$. 

(ii) Often signal models are described in the above continuous-time setting but sampled measurements $z_k$ of $z(t)$ are available. Write down a hybrid continuous-discrete version of the EKF in corrector-predictor form.

Problem 5. Consider a pendulum of length $\ell$ that subtends an angle $\theta(t)$ with a vertical line through its pivot. The pendulum’s angular acceleration and measurements of its instantaneous horizontal position (from the vertical) may be modelled as $\frac{d^2\theta(t)}{dt^2} = -\frac{g}{\ell}\sin(\theta(t)) + w(t)$ and $z(t) = \ell\sin(\theta(t)) + v(t)$, respectively, where $g$ is the gravitational constant, $w(t)$ and $v(t)$ are stochastic inputs.

(i) Set out the pendulum’s equations of motion in a state-space form and write down the continuous-time EKF for estimating $\theta(t)$ from $v(t)$.

(ii) Use Euler’s first-order integration formula to discretise the above model and then detail the corresponding discrete-time EKF.

10.9 Glossary

$\nabla f$ The gradient of a function $f$, which is a row-vector of partial derivatives.

$\nabla^T\nabla f$ The Hessian of a function $f$, which is a matrix of partial derivatives.

$\text{tr}(P_i)$ The trace of a matrix $P_i$, which is the sum of its diagonal terms.

FM Frequency modulation.

$\Delta f$ The forward difference operator with $\Delta f_i = f_i^{(1)} - f_{i-1}^{(1)}$.

10.10 References


“What we observe is not nature itself, but nature exposed to our mode of questioning.” Werner Heisenberg


“We know nothing in reality; for truth lies in an abyss.” Democritus

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This book describes the classical smoothing, filtering and prediction techniques together with some more recently developed embellishments for improving performance within applications. It aims to present the subject in an accessible way, so that it can serve as a practical guide for undergraduates and newcomers to the field. The material is organised as a ten-lecture course. The foundations are laid in Chapters 1 and 2, which explain minimum-mean-square-error solution construction and asymptotic behaviour. Chapters 3 and 4 introduce continuous-time and discrete-time minimum-variance filtering. Generalisations for missing data, deterministic inputs, correlated noises, direct feedthrough terms, output estimation and equalisation are described. Chapter 5 simplifies the minimum-variance filtering results for steady-state problems. Observability, Riccati equation solution convergence, asymptotic stability and Wiener filter equivalence are discussed. Chapters 6 and 7 cover the subject of continuous-time and discrete-time smoothing. The main fixed-lag, fixed-point and fixed-interval smoother results are derived. It is shown that the minimum-variance fixed-interval smoother attains the best performance. Chapter 8 attends to parameter estimation. As the above-mentioned approaches all rely on knowledge of the underlying model parameters, maximum-likelihood techniques within expectation-maximisation algorithms for joint state and parameter estimation are described. Chapter 9 is concerned with robust techniques that accommodate uncertainties within problem specifications. An extra term within Riccati equations enables designers to trade-off average error and peak error performance. Chapter 10 rounds off the course by applying the afore-mentioned linear techniques to nonlinear estimation problems. It is demonstrated that step-wise linearisations can be used within predictors, filters and smoothers, albeit by forsaking optimal performance guarantees.

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